



The Lower Bound Estimation for the Number of Zeros of Random Transcendental Polynomial

Research Article

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Abstract: The object of this paper is to find a lower bound estimation for the number of zeros of the random transcendental equation $\sum_{v=0}^n d_v \xi_v(\omega) z^v = 0$, subject to the condition that the coefficients are non-identically distributed dependent random variables. Throughout the paper n is considered to be very large and μ 's denote positive constants assuming different values in different occurrences.

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1. Introduction

$N_n(\omega)$ is the number of zeros of the random transcendental Polynomial

$$f_n(z, \omega) = \sum_{v=0}^n d_v \xi_v(\omega) z^v \quad (1)$$

and d_v 's be non-zero real numbers, when $\xi_v(\omega)$'s are symmetric stable variates with characteristic function

$$\exp(-C|t|^\alpha), \quad C > 0, \quad 1 < \alpha \leq 2.$$

Assuming the coefficients $\xi_v(\omega)$'s are non-identically distributed dependent random variables on probability space (Ω, B, P) .

Define Normal Distribution with mean zero and joint density function

$$M^{\frac{1}{2}} (2\pi)^{-\frac{n}{2}} \exp\left(-\frac{1}{2} \bar{a}' M \bar{a}\right), \quad \bar{a}' = [\xi_1(\omega), \xi_2(\omega), \dots, \xi_n(\omega)] \quad (2)$$

where M^{-1} is the moment matrix with $\sigma_i = 1$, $\rho_{ij} = \rho$, $0 < \rho < 1$, $i \neq j$, $i, j = 0, 1, \dots, n$ and \bar{a} is the column vector whose transpose is \bar{a}' follows from [1], [2], [3]. Let G be the exceptional set defined by $G = \{\omega | N_n(\omega) > \mu (\log \log n)^2 \log n\}$, where μ is a positive constant. We introduce a notation $\lambda = \log n$ and M be the integer defined by $M = \left[\alpha \lambda \frac{k_n}{t_n} \right] + 1$, where α is a positive constant and $[.]$ implies the greatest integer function. Let k be the integer determined by $M^{2k} \leq n < M^{2k+2}$.

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Again we introduce a new notation $\lambda_m = m^{\frac{1}{2}} \log n$ the random algebraic polynomial and M_m be a sequence of integers defined by $M_m = \left[b \left(\frac{k_n}{t_n} \right) \log n \right] + 1$, $m = 1, 2, \dots$, where b is a positive constant. Let k be the integer determined by

$$(2k)! M_n^{2k} \leq n < (2k+2)! M_n^{2k+2}.$$

We shall use the fact that each $\xi_v(\omega)$ has marginal frequency function $\frac{1}{2\pi} \exp\left(-\frac{\omega^2}{2}\right)$. In this paper we have established two theorems where the second theorem is a modified and more effective interpretation of the first one in the sense of Evans [1], since the exceptional set G obtained in this case is independent of n . Throughout this paper n is considered to be very large and μ 's, α 's, b 's denote positive constants assuming different values in different occurrences.

2. Preliminary and Some Results

Theorem 2.1. *Let $f_n(z, \omega) = \sum_{v=0}^n d_v \xi_v(\omega) z^v$ be a random transcendental polynomial, where the $\xi_v(\omega)$'s are non-identically distributed dependent random variables with mean zero and joint density function given by (2). Let d_v 's be non-zero real numbers such that*

$$\frac{k_n}{t_n} = O(\log n), \text{ where } k_n = \max_{0 \leq v \leq n} |d_v|, t_n = \min_{0 \leq v \leq n} |d_v|$$

Then there exist a positive integer n_0 Such that for $n > n_0$

$$N_n(\omega) \geq \mu \frac{\log n}{\log \left(\frac{k_n}{t_n} \right) \log n}$$

and

$$P(G) \leq \mu' \frac{\log \left(\frac{k_n}{t_n} \right) \log n}{\log n}.$$

Where $N_n(\omega)$ is the number of real zeros of the polynomial and G is the exceptional set by $G = \{\omega | N_n(\omega) > \mu (\log \log n)^2 \log n\}$ and $P(G) \rightarrow 0$ as $n \rightarrow \infty$.

Proof. Let

$$\lambda = \log n \tag{3}$$

and M be the integer defined by

$$M = \left[\alpha \lambda \frac{k_n}{t_n} \right] + 1 \tag{4}$$

where α is a positive constant and $[.]$ implies the greatest integer function. Let k be the integer determined by

$$M^{2k} \leq n < M^{2k+2} \tag{5}$$

It follows from (3), (4), and (5) that, for two constants μ_1 and μ_2 ,

$$\mu_1 \frac{\log n}{\log \left(\frac{k_n}{t_n} \right) \log n} \leq k \leq \mu_2 \frac{\log n}{\log \left(\frac{k_n}{t_n} \right) \log n} \tag{6}$$

We shall consider $f_n(z, \omega)$ at the points

$$z_m = (1 - M^{-2m})^{\frac{1}{2}} \tag{7}$$

for $m = \lfloor \frac{k}{2} \rfloor + 1, \lfloor \frac{k}{2} \rfloor + 2, \dots, k$. Let

$$f_n(z_m, \omega) = \sum_1 d_v \xi_v(\omega) z_m^v + \left(\sum_2 + \sum_3 \right) d_v \xi_v(\omega) z_m^v = A_m(\omega) + R_m(\omega)$$

where v ranges from $M^{2m-1} + 1$ to M^{2m+1} in \sum_1 and from 0 to M^{2m-1} in \sum_2 and from $M^{2m+1} + 1$ to n in \sum_3 where

$$A_m(\omega) = \sum_1 d_v \xi_v(\omega) z_m^v \quad \text{and}$$

$$R_m(\omega) = \left(\sum_2 + \sum_3 \right) d_v \xi_v(\omega) z_m^v$$

The following lemmas are necessary for the rest proof of Theorem 2.1

Lemma 2.2. For $\alpha_1 > 0$, $\sigma_m > \alpha_1 t_n M^{2m}$, where

$$\sigma_m^2 = (1-\rho) \sum_1 d_v^2 z_m^{2v} + \rho \left(\sum_1 d_v z_m^v \right)^2, \quad (0 < \rho < 1) \quad (*)$$

Proof.

$$\sum_1 d_v z_m^v > t_n \sum_1 z_m^v > t_n M^{2m} \left(\frac{B}{A\sqrt{e}} \right) \Rightarrow \left(\sum_1 d_v z_m^v \right)^2 > \alpha_1^2 t_n^2 M^{4m}, \quad (i)$$

α_1 is a positive constant. Where A, B and constants satisfying the relations, $A > 1$ and $0 < B < 1$. Again

$$\sum_1 d_v^2 z_m^{2v} > M^{2m} t_n^2 \left(\frac{B}{Ae} \right) \quad (ii)$$

From (i) and (ii), (*) becomes $\sigma_m^2 > \alpha_1^2 t_n^2 M^{4m}$ (where α_1 is a positive constant) and hence $\sigma_m > \alpha_1 t_n M^{2m}$. Which gives the result.

Lemma 2.3.

$$P \left\{ \omega : \left| \sum_2 d_v \xi_v(\omega) z_m^v \right| > \lambda \tilde{\sigma}_m \right\} < \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\lambda^2}{2}}}{\lambda} \quad \text{where } \tilde{\sigma}_m^2 = (1-\rho) \sum_2 d_v^2 z_m^{2v} + \rho \left(\sum_2 d_v z_m^v \right)^2, \quad (0 < \rho < 1).$$

Proof. Let $F(z)$ be the distribution function of $\sum_2 d_v \xi_v(\omega) z_m^v$. Then

$$\begin{aligned} P \left\{ \omega : \left| \sum_2 d_v \xi_v(\omega) z_m^v \right| > \lambda \tilde{\sigma}_m \right\} &= 1 - \{F(\lambda \tilde{\sigma}_m) - F(-\lambda \tilde{\sigma}_m)\} \\ &= \sqrt{\frac{2}{\pi}} \int_{\lambda}^{\infty} e^{-\frac{t^2}{2}} dt < \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\lambda^2}{2}}}{\lambda} \end{aligned}$$

Lemma 2.4.

$$P \left\{ \omega : \left| \sum_3 d_v \xi_v(\omega) z_m^v \right| > \lambda \tilde{\tilde{\sigma}}_m \right\} < \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\lambda^2}{2}}}{\lambda}, \quad \text{where } \tilde{\tilde{\sigma}}_m^2 = (1-\rho) \sum_3 d_v^2 z_m^{2v} + \rho \left(\sum_3 d_v z_m^v \right)^2, \quad (0 < \rho < 1)$$

Lemma 2.5. For a fixed m ,

$$P \{ \omega : |R_m(\omega)| < \sigma_m \} > 1 - 2\sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\lambda^2}{2}}}{\lambda}$$

Proof. For given m , we have $|R_m(\omega)| < \lambda(\tilde{\sigma}_m + \tilde{\tilde{\sigma}}_m)$ again

$$\sum_2 d_v^2 z_m^{2v} \leq 2k_n^2 M^{2m-1} \text{ and } \sum_2 d_v z_m^v \leq 2k_n M^{2m-1}$$

Hence $\tilde{\sigma}_m^2 \leq \alpha_2^2 k_n^2 M^{2m-1} < \alpha_2^2 k_n^2 M^{4m-2}$, $\alpha_2 > 0$. Similarly $\tilde{\tilde{\sigma}}_m^2 \leq \alpha_3^2 k_n^2 M^{2m-1} < \alpha_3^2 k_n^2 M^{4m-2}$, $\alpha_3 > 0$. Thus

$$\begin{aligned} |R_m(\omega)| &< \lambda(\alpha_2 + \alpha_3) k_n M^{2m-1} \\ &< \frac{\left\{ \lambda \left(\frac{\alpha_2 + \alpha_3}{\alpha_1} \right) \frac{k_n}{t_n} \sigma_m \right\}}{M} \text{ by Lemma 2.1} \\ &< \sigma_m \text{ by the definition of } M \end{aligned}$$

Since the distribution function of $A_m(\omega)$ is

$$\frac{1}{\sqrt{2\pi} \sigma_m} \int_{-\infty}^z \exp \left[- \left(\frac{t^2}{2\sigma_m^2} \right) \right] dt$$

where

$$\sigma_m^2 = (1 - \rho) \sum_1 d_v^2 z_m^{2v} + \rho \left(\sum_1 d_v z_m^v \right)^2, \quad 0 < \rho < 1$$

The distribution function of $\frac{A_m}{\sigma_m}$ is

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^z \exp \left[- \left(\frac{t^2}{2} \right) \right] dt = D(z) \text{ (say)}$$

Now let us define random events E_m and F_m by

$$\begin{aligned} E_m &= \{ \omega : A_{2m}(\omega) \geq \sigma_{2m}, A_{2m+1}(\omega) < -\sigma_{2m-1} \} \\ F_m &= \{ \omega : A_{2m}(\omega) < -\sigma_{2m}, A_{2m+1}(\omega) \geq \sigma_{2m+1} \} \end{aligned}$$

It can be easily seen that $P(E_m \cup F_m) > \delta > 0$, where δ is a positive constant. Then proceeding exactly as Samal and Mishra [3] we shall get the following results.

$$N_n(\omega) > \mu_1 k \geq \frac{\mu \log n}{\log \left(\frac{k_n}{t_n} \right) \log n} \quad \text{by (6)}$$

and

$$P(G) < \mu_1 k \frac{1}{\lambda e^{\frac{\lambda^2}{2}}} + \frac{\mu_2}{k} \leq \mu' \left\{ \frac{\log \left(\frac{k_n}{t_n} \right) \log n}{\log n} \right\}$$

Since $\frac{k_n}{t_n} = 0(\log n)$, $P(G) \rightarrow 0$ as $n \rightarrow \infty$. Hence the Theorem 2.1. \square

Theorem 2.6. *Let*

$$f_n(z, \omega) = \sum_{v=0}^n d_v \xi_v(\omega) z^v$$

be a random transcendental polynomial, where the $\xi_v(\omega)$'s are non-identically distributed dependent random variables with mean zero and joint density function given by (2). Let d_v 's be non-zero real numbers, such that

$$\frac{k_n}{t_n} = 0(\log n), \text{ where } k_n = \max_{0 \leq v \leq n} |d_v|, t_n = \min_{0 \leq v \leq n} |d_v|$$

Then there exist a positive integer n_0 such that for $n > n_0$

$$N_n(\omega) \geq \mu \frac{\log n}{\log \left(\frac{k_n}{t_n} \log n \right)} \quad \text{and}$$

$$P(G) \leq \mu' \left\{ \frac{\log \left(\log \left(\frac{k_{n_0}}{t_{n_0}} \right) \log n_0 \right)}{\log n_0} \right\}^{\frac{1}{2}}$$

$N_n(\omega)$ is the number of real zeros of the polynomial and G is the exceptional set.

Proof. Let

$$\lambda_m = m^{\frac{1}{2}} \log n \quad (8)$$

and M_m be a sequence of integers defined by

$$M_m = \left[b \left(\frac{k_n}{t_n} \right) \log n \right] + 1, \quad m = 1, 2, \dots \quad (9)$$

where b is a positive constant. Let k be the integer determined by

$$(2k)! M_n^{2k} \leq n < (2k+2)! M_n^{2k+2} \quad (10)$$

It follows from (9) and (10) that for two constants μ_1 and μ_2

$$\mu_1 \frac{\log n}{\log \left(\frac{k_n}{t_n} \log n \right)} \leq k \leq \mu_2 \frac{\log n}{\log \left(\frac{k_n}{t_n} \log n \right)} \quad (11)$$

We consider $f_n(z, \omega)$ at the points

$$z_m = \left\{ 1 - \frac{1}{(2k)! M_m^{2k}} \right\}^{\frac{1}{2}} \quad (12)$$

for $m = \left[\frac{k}{2} \right] + 1, \left[\frac{k}{2} \right] + 2, \dots, k$. for large n we write

$$f_n(z_m, \omega) = A_m(\omega) + R_m(\omega)$$

where

$$A_m(\omega) = \sum_1 d_v \xi_v(\omega) z_m^v$$

$$\text{and} \quad R_m(\omega) = \left(\sum_2 + \sum_3 \right) d_v \xi_v(\omega) z_m^v$$

and the index v ranges from $(2m-1)! M_m^{2m-1} + 1$ to $(2m+1)! M_m^{2m+1}$ in \sum_1 and from 0 to $(2m-1)! M_m^{2m-1}$ in \sum_2 and from $(2m+1)! M_m^{2m+1} + 1$ to n in \sum_3

The following lemmas are necessary for the rest proof of Theorem 2.6.

Lemma 2.7. For $r_1 > 0$, $\sigma_m > r_1 t_n (2m)! M_m^{2m}$, where

$$\sigma_m^2 = (1-\rho) \sum_1 d_v^2 z_m^{2v} + \rho \left(\sum_1 d_v z_m^v \right)^2, \quad 0 < \rho < 1 \quad (*)$$

Proof.

$$\sum_1 d_v^2 z_m^{2v} > t_n^2 \left((2m)! M_m^{2m} \frac{B}{Ae} \right)^2, \quad \text{where } A > 1, 0 < B < 1 \quad (i)$$

and

$$\left(\sum_1 d_v z_m^v \right)^2 > t_n^2 \left(\sum_1 z_m^v \right)^2 > t_n^2 \left((2m)! M_m^{2m} \frac{B}{A\sqrt{e}} \right)^2 \quad (ii)$$

From (i) and (ii), (*) becomes

$$\sigma_m^2 > (1-\rho) t_n^2 \left((2m)! M_m^{2m} \frac{B}{Ae} \right)^2 + \rho t_n^2 \left((2m)! M_m^{2m} \frac{B}{A\sqrt{e}} \right)^2$$

And hence

$$\sigma_m > r_1 t_n (2m)! M_m^{2m}$$

Again the following two Lemmas 2.8 and 2.9 can be proved as the process that has been adopted in Theorem 2.1.

Lemma 2.8.

$$P \left\{ \omega: \left| \sum_2 d_v \xi_v(\omega) z_m^v \right| < \lambda q_m \right\} < \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\lambda^2}{2}}}{\lambda} \quad \text{where } q_m^2 = (1-\rho) \sum_2 d_v^2 z_m^{2v} + \rho \left(\sum_2 d_v z_m^v \right)^2.$$

Lemma 2.9.

$$P \left\{ \omega: \left| \sum_3 d_v \xi_v(\omega) z_m^v \right| < \lambda Q_m \right\} < \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\lambda^2}{2}}}{\lambda} \quad \text{where } Q_m^2 = (1-\rho) \sum_3 d_v^2 z_m^{2v} + \rho \left(\sum_3 d_v z_m^v \right)^2.$$

Lemma 2.10. For a fixed m , $P \{ \omega: |R_m(\omega)| < \sigma_m \} > 1 - 2\sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\lambda^2}{2}}}{\lambda}$.

Proof. For a given m , we have $|R_m(\omega)| < \lambda(q_m + Q_m)$. Now

$$\begin{aligned} \sum_2 d_v^2 z_m^{2v} &\leq 2k_n^2 (2m-1)! M_m^{2m-1} \quad \text{and} \\ \sum_2 d_v z_m^v &\leq 2k_n (2m-1)! M_m^{2m-1} \end{aligned}$$

$q_m^2 \leq r_2 k_n (2m-1)! M_m^{2m-1}$, where r_2 is a positive constant $< r_2 k_n (2m)! M_m^{2m-1}$. Similarly $Q_m^2 < r_3 k_n (2m)! M_m^{2m-1}$, where r_3 is a positive constant. So,

$$|R_m(\omega)| < \lambda k_n \frac{M_m^{2m} (r_2 + r_3)}{M_m} < \sigma_m.$$

Therefore $|R_m(\omega)| < \sigma_m$ except for a set of measure at most

$$2\sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\lambda^2}{2}}}{\lambda} \quad (\text{by definition of } M_m)$$

$$\text{for } m = \left[\frac{k}{2} \right] + 1, \left[\frac{k}{2} \right] + 2, \dots, k.$$

Now defining the E_m and F_m as in Theorem 2.1, we can have $P(E_m \cup F_m) > \delta > 0$, where δ is an absolute constant. Now let us define random variables $X_m(\omega)$, $Y_m(\omega)$ and net G_m as

$$X_m(\omega) = \begin{cases} 1, & \text{if } \omega \in E_m \cup F_m \\ 0, & \text{if } \omega \in (E_m \cup F_m)'. \end{cases}$$

Thus

$$P \{ \omega: X_m(\omega) = 1 \} = \delta$$

and

$$P \{ \omega: X_m(\omega) = 0 \} = 1 - \delta$$

Let $G_m = \{ \omega: |R_{2m}(\omega)| < \sigma_{2m} \text{ and } R_{2m+1}(\omega) < \sigma_{2m+1} \}$ and

$$Y_m(\omega) = \begin{cases} 0, & \text{if } \omega \in G_m; \\ 1, & \text{if } \omega \in (G_m)'. \end{cases}$$

Let $T_m(\omega) = X_m(\omega) - X_m(\omega) Y_m(\omega)$. If $T_m(\omega) = 1$, then there exist a zero of the polynomial in the interval (z_{2m}, z_{2m+1}) .

Now proceeding as Samal and Mishra [3] we get,

$$\begin{aligned} N_n(\omega) &\geq \mu_1 k \\ &\geq \mu \frac{\log n}{\log \left(\frac{k_n}{t_n} \log n \right)} && \text{by (11)} \\ \text{and } P(G) &\leq \frac{\mu_2}{k_0} + \mu_3 \sum_{k \geq 2k_0-1} \frac{\exp \left(-\frac{\lambda_{m_0}^2}{2} \right)}{\lambda_{m_0}} \\ &\leq \frac{\mu_2}{k_0} + \mu_4 \sum_{k \geq 2k_0-1} \frac{1}{\lambda_{m_0}^3} \\ &\leq \frac{\mu_2}{k_0} + 2\mu_4 \sum_{k \geq k_0} \frac{1}{k^{\frac{3}{2}}} \\ &\leq \frac{\mu_2}{k_0} + 2\mu_4 \left(\frac{2}{k_0^{\frac{1}{2}}} \right) \\ &\leq \mu \left\{ \frac{\log \left(\log \left(\frac{k_{n_0}}{t_{n_0}} \right) \log n_0 \right)}{\log n_0} \right\}^{\frac{1}{2}} && \text{by (11)} \end{aligned}$$

hence the Theorem 2.6. □

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